Diagnosis of UAV Pitot Tube Defects Using Statistical Change Detection

Hansen, Søren; Blanke, Mogens; Adrian, Jens

Published in:
7. Symposium on Intelligent Autonomous Vehicles

Publication date:
2010

Document Version
Early version, also known as pre-print

Link back to DTU Orbit

Citation (APA):
Diagnosis of UAV Pitot Tube Failure Using Statistical Change Detection

Søren Hansen *,** Mogens Blanke *,** Jens Adrian **

* Technical University of Denmark, Department of Electrical Engineering, Automation and Control Group, Elektrovej B. 326, DK-2800 Lyngby. Denmark.
** Danish Forces Joint UAV Team, Naval Weapons School, Gnibenvej 55, DK-4583 Sjællands Odde. Denmark.
*** CeSOS, Marine Technology Centre, Norwegian University of Science and Technology, NTNU, NO-7491 Trondheim, Norway.
(E-mail: sh@elektro.dtu.dk, mb@elektro.dtu.dk, adrian@mil.dk)

Abstract: Unmanned Aerial Vehicles need a large degree of tolerance to faults. One of the most important steps towards this is the ability to detect and isolate faults in sensors and actuators in real time and make remedial actions to avoid that faults develop to failure. This paper analyses the possibilities of detecting faults in the pitot tube of a small unmanned aerial vehicle, a fault that easily causes a crash if not diagnosed and handled in time. Using as redundant information the velocity measured from an onboard GPS receiver, the air-speed estimated from engine throttle and the pitot tube based airspeed, the paper analyses the properties of residuals. A dedicated change detector is suggested that works on pre-whitened residuals and a generalised likelihood ratio test is derived for a Cauchy probability density, which the residuals are observed to have. A detection scheme is obtained using a threshold that provides desired quantities of false alarm and detection probabilities. Fault detectors are build based on raw residual data and on a whitened edition of these. The two detectors are compared against recorded telemetry data of an actual event where a pitot tube defect occurred.

Keywords: Fault detection, Change detection, Unmanned Aerial Vehicle, Pitot tube

1. INTRODUCTION

Defects on sensors can have catastrophic consequences for airplanes, specially smaller airplanes, which do not have the same sensor redundancy that is available on larger aircraft. It is therefore important to be able to detect whether a sensor defect has occurred. One of the most important sensors is the pitot tube that measures airspeed of the vehicle. This sensor is very exposed because of its position in the airstream and can easy be clogged by dust or water particles that freeze at higher altitudes, when they come in contact with the airplane body.

The solution to these clogging problems usually employed on larger aircrafts is to install several pitot tubes with build in heating devices to have a redundant system that can accommodate icing. Because of weight and space limitations, adding more sensors is usually not an option on smaller unmanned aerial vehicles (UAV). Therefore, a different approach must be taken to diagnose and accommodate faults. One way could be to have artifact readings detected and replaced with estimated values. Detection of faults and fault-tolerance for UAVs has a lot of focus and, as described in Ducard (2009), many parts of the aircraft control and operation can benefit from using fault tolerant methods. A systematic approach to fault detection is described in Fravolini et al. (2009) and some of the applications of these methods are, detection of mechanical defects, like stuck control surfaces. These were studied in Bateman et al. (2009) and Park et al. (2009) where active methods were used to isolate faults. Observer based fault diagnosis was investigated in e.g. Heredia et al. (2008). Nonlinear models that describe the aircraft can also be used in fault diagnosis, this was demonstrated on small helicopters in Freddi et al. (2009).

This paper investigates the use of GPS velocity measurements and propeller thrust readings to generate speed information that is redundant compared to the pitot tube airspeed. Residual values are formed from which defects on the pitot tube can be isolated. The detection is done using statistical information gathered from telemetry records from a Banshee UAV (see fig. 1). Advantage is taken of availability of real data to derive probability density functions and establish spectral properties of data. Dedicated change detection is derived based on observed signal properties and detectors are convincingly tested on real data from an incident involving icing of the pitot tube.

2. PLATFORM

The UAV studied in this paper is a Banshee target drone build by the British company Meggitt Defence System.

The airplane is controlled by an autopilot and the operator sends set-point requests to height, speed and heading. Telemetry data are sent from the plane and logged at a ground station. These data are the basis for the investigations presented in this paper.
3. MODEL

The airplane is modelled by the dynamic and kinematic equations, which describe its motion through the air. In order to describe these concepts mathematically, a number of reference frames are needed. These are the ECEF (Earth centred fixed) frame, the VCE (Vehicle carried earth) frame, the body frame and the wind frame, see e.g Stevens and Lewis (2003) for details.

3.1 Airplane motion parameters

The airplane’s motion is described by its velocities in the body coordinate system

$$\mathbf{v}_b = [u, v, w]^T,$$

and the rotational rate about the body axes,

$$\omega = [p, q, r]^T.$$ (2)

The Euler angles used to describe the rotation between the vehicle carried coordinate frame and the body frame of the aircraft are,

$$\Phi = [\phi, \theta, \psi]^T.$$ (3)

This is also the attitude and heading information.

3.2 Kinematics

The kinematics of the airplane relates the angular velocities $\omega$ to the rate of change in the Euler angles $\Phi$, which are used in the rotation between body and the VCE coordinate system. This is given by the following relation

$$\dot{\Phi} = L(\Phi)\omega$$ (4)

$$L(\Phi) = \begin{bmatrix}
1 & \tan \theta \sin \phi & \tan \theta \cos \phi \\
0 & \cos \phi & -\sin \phi \\
0 & \sin \phi / \cos \theta \cos \phi / \cos \theta
\end{bmatrix}$$ (5)

The change in position in the ECEF reference frame can be described by rotating the translational velocities $\mathbf{v}_b$ to the navigation frame. The transformation matrix involved in this depends on $\Phi$.

3.3 Dynamics

The airplane is seen as a rigid body moving through the air. Newton’s and Euler’s laws for linear and angular momentum relate forces and moments acting on the rigid body to linear and angular velocities and accelerations. Since the UAV is flying at low speed and height over a relatively small area, we assume that the Earth is flat, and Coriolis forces and the centripetal force from the curvature of the Earth can be neglected.

The airplane is affected by a thrust from the engine $\mathbf{F}_T$, an aerodynamic force $\mathbf{F}_A$ arising from the lift and drag from the airplane body and, naturally, the gravitational force from Earth $\mathbf{F}_G$. With mass $m$,

$$\frac{d}{dt}(mv_b) = \sum \mathbf{F} = \mathbf{F}_T + \mathbf{F}_A + \mathbf{F}_G$$ (6)

Since this equation is expressed in the inertial frame (the Earth system) the equation of Coriolis is used to find the derivative,

$$mv_b + \omega \times mv_b = \mathbf{F}_T + \mathbf{F}_A + \mathbf{F}_G$$ (7)

For an in depth description of airplane dynamics see for instance Stevens and Lewis (2003).

3.4 Thrust

The Banshee’s thrust is delivered by a rear mounted internal combustion engine that drives a 2 bladed wood propeller. The actual force delivered depends on the physical dimensions of the propeller, the engines revolutions and the density of the air. Given a propeller with diameter $d_p$ rotating at $n$ revolutions per second the thrust force is given by:

$$F_T = k_T \cdot \rho n^2 d_p^4$$ (8)

where $k_T$ is the thrust coefficient and $\rho$ is the density of air. Wind tunnel tests (see Lesley (1939)) of propellers of the used type show that the thrust coefficient is approximately linear with the advance ratio $J$ as seen in Lesley (1939) figure 5. The advance ratio is given by:

$$J = \frac{v}{nd_p}.$$ (9)

In nominal flight the advancement ratio is in the range 0.9 to 1.1.

Propellers are usually characterised by their diameter $d_p$ and pitch $P_p$. These values can be combined to a non-dimensional pitch given by:

$$P' = \frac{P_p}{d_p}$$ (10)

In Nissen (2002) the following scaling for the thrust coefficient was proposed.

$$k_T(J) \simeq k_{T,ref}(J - (P' - P'_{ref}))$$ (11)

where $k_{T,ref}$ and $P'_{ref}$ is the thrust coefficient and non-dimensional pitch of the reference signal respectively. The reference signals originates from different propellers tested in Lesley (1939). Since the $26 \times 30$ inch propeller used on the Banshee has $P'$, which is equal to one of the tested propellers, the characteristics of this propeller is used from

![Fig. 1. Banshee UAV ready for launch at the Danish Naval Weapons School at Sjælland's Odde. (Photo: VFD)](image-url)
are crucial for the correctness of these conversions. These
estimates show nonlinear observer designs for thrust
et al. (2009) showed nonlinear observer designs for thrust
(1998) applied an adaptive observer scheme and Pivano
pose of fault diagnosis, more elaborate and sophisticated
since the angle of attack
α is usually small this force primarily consists of the drag
on the airplane. This is modelled on basis of recorded data
to vary with \( \theta \). It is assumed that the engine is aligned
with the airplane body and therefore \( F_{Tx} = F_T \). In straight
and level flight with no accelerations present we can estimate
the forward air speed of the plane by inserting the thrust
expression from equation 13.
\[
\mathbf{\dot{u}} = -\frac{T_{nn}n^2 - F_{Ax} - mg \sin \theta}{T_{nu}n}
\]  
While this simple estimate was found to suffice for the pur-
purpose of fault diagnosis, more elaborate and sophisticated
nonlinear observers and estimators have been suggested in
the literature. Zhou and Blanke (1989) described a way to
estimate state and parameters in nonlinear systems with a
structure similar to the thrust equation here, Blanke et al.
(1998) applied an adaptive observer scheme and Pivano
et al. (2009) showed nonlinear observer designs for thrust
estimation.

4. SIGNAL ANALYSIS

The airspeed of the UAV is measured by pressure differ-
ences between the pitot tube input and the static vent input.
This gives the difference between the stagnation pres-
sure and the static pressure which, according to Bernoulli’s
theorem, is the dynamic pressure \( \bar{q} \),
\[
\bar{q} = \frac{1}{2} \rho \bar{V}^2
\]  
From this the planes airspeed \( V \) can be calculated.

In order to detect an error in the airspeed indicator this
sensor value must be compared to similar measurements.
The GPS receiver gives a speed from its internal Kalman
filter. This corresponds to the speed in an ECEF reference
frame, but using knowledge about the wind speed and
direction it can be transformed to airspeed measured in
the pitot tube.

As seen in the previous chapter the speed in the body
\( x \)-direction can be derived from the thrust equations
developed by the engine, by using equation 17. This
measurment can also be converted to match the airspeed
measured by the pitot tube system.

The estimation of the wind direction and wind speed
are crucial for the correctness of these conversions. These
values are also hard to estimate with common means.
However the wind is measured at the ground station, and
also estimated during takeoff where the airplane always
climbs into the wind direction. These estimates prove to
give good results in practice.

The simplest way of comparing the three airspeed
measures are by subtracting them from each other. This
creates three residual signals which in the nominal case
should have a values around 0. If we denote the airspeed
measured by the pitot tube \( v_{airspeed} \), the airspeed calcu-
lated from the GPS \( v_{gps} \) and the airspeed calculated from
the engine velocity \( v_{thrust} \) the following parity matrix can
be formed:

\[
\begin{bmatrix}
R_1 \\
R_2 \\
R_3
\end{bmatrix} =
\begin{bmatrix}
v_{airspeed} & v_{gps} & v_{thrust}
\end{bmatrix}
\]

An 's' in list 19 means that errors on this value is strongly
detectable by the given residual. All the residual are
generated from the difference between measurements of
the airspeed velocity. This means that if the residual is
different from zero it indicates a fault, on one of the
signals. Therefore an detection of a DC-signal of unknown
magnitude, different from zero, in the residual, would
indicate an error.

Since only \( R_1 \) and \( R_2 \) contains pitot pressure information
only two of these residuals are analysed onwards, however
the same methods are applied to \( R_3 \) to obtain fault isolation.

Fig. 2. Time development and histogram for residual \( R_1 \)
and \( R_2 \) in the fault free case.

Figure 2 shows a time history and a histogram for \( R_1 \)
and \( R_2 \) in the fault free case.

As seen from the power spectrum densities in figure 3
the noise present at the residual signals are not white.
Since white noise is one of the requirements for most
statistical change detectors to perform optimal, the low-
pass filtered nature of the noise should be removed. One
solution of dealing with coloured noise is to filter white
noise through a suitable filter function to take into account
the correlations present in the coloured noise. This can be
created from a large record of data where all the signals
A filter function $H$ which converts a white noise input to a coloured noise output according to

$$S_{rr}(\omega) = |H(\omega)|^2 S_{ww}(\omega) = H(j\omega)H^*(j\omega)S_{ww}(\omega) \quad (20)$$

can be created. The function should satisfy the shape given in figure 3 and satisfy that

$$\frac{1}{2\pi} \int_{-\infty}^{\infty} S_{rr}(\omega) d\omega = \sigma_r^2 \quad (21)$$

where $\sigma_r^2$ is the variance of the residual noise. In order to satisfy the shape in the following PSD function is chosen.

$$P(\omega) = K \frac{\tau_1 \omega^2 + 1}{(\tau_1 \omega^2 + 1)^2} \frac{\tau_2 \omega^2 + 1}{(\tau_2 \omega^2 + 1)^2} \quad (22)$$

where $\tau_1$ is the reciprocal cut-off frequency between the 0 dB/decade and the $-40$ dB/decade parts and $\tau_2$ is the time constant between $-40$ dB/decade and $-80$ dB/decade parts, and $K$ is the magnitude. This corresponds to the following filter function.

$$P(\omega) = H(j\omega)H^*(j\omega) = \quad (23)$$

$$\sqrt{K} \frac{\sqrt{\tau_1} j\omega + 1}{(\sqrt{\tau_1} j\omega + 1)^2} \frac{\sqrt{\tau_2} j\omega + 1}{(\sqrt{\tau_2} j\omega + 1)^2} \times$$

$$\sqrt{K} \frac{-\sqrt{\tau_1} j\omega + 1}{(-\sqrt{\tau_1} j\omega + 1)^2} \frac{-\sqrt{\tau_2} j\omega + 1}{(-\sqrt{\tau_2} j\omega + 1)^2}$$

The constants in the equation are found by optimising $H(j\omega)$ to fit the actual spectra.

Expressing the filter function in terms of $s$ gives

$$H(s) = \sqrt{K} \frac{\sqrt{\tau_1} s + 1}{(\sqrt{\tau_1} s + 1)^2} \frac{\sqrt{\tau_2} s + 1}{(\sqrt{\tau_2} s + 1)^2} \quad (24)$$

The filter is implemented digitally using an IIR representation. Since filtering white noise through $H(s)$ should create a signal with a flat power spectrum similar to the one in figure 3 filtering $R_1$ and $R_2$ through $H^{-1}(s)$ should create a signal with a flat power spectrum. Figure 5 shows the result of the whitening. Both residuals features now white noise. In figure 4 the time development and the histogram is plotted. Neither of the residuals are now Gaussian distributed. Instead they follow the Cauchy distribution (equation 34) very well as indicated by the green dotted line in the histogram. The Gaussian nature that the residuals has before whitening comes from the effect of the low-pass filtering, which creates the power spectra observed. When this is removed during the whitening process the Gaussianity is also removed.

5. DETECTORS

Two approaches to change detection is investigated. In the first one the raw residual signals are used and the next one the whitened signals are used.
5.1 Raw residual detector

If the vocabulary from statistical change detection is used the scenario that it is desirable to analyse is: Detect an unknown DC-level in Gaussian white noise. If the signal is denoted \(x\) we wish to distinguish between the following two hypotheses:

\[
\mathcal{H}_0 : x[n] = w[n] \\
\mathcal{H}_1 : x[n] = A + w[n]
\]  

(25)

The unknown parameters are: The DC-level \(A\) and the variance of the noise \(\sigma_w^2\). The GLRT (Generalised Likelihood Ratio Test) for a linear model formulation can be used to solve such a problem. However, GLRT algorithm assumes that the signal noise is white. As emphasised, this is not the case for the raw residual. However, the algorithm is tested anyway using the raw residual, to see how it performs. The general GLRT detector is given by

\[
L_G(x) = \frac{p(\mathbf{x}; \hat{\Theta}_1)}{p(\mathbf{x}; \hat{\Theta}_0)} > \gamma
\]

(26)

If the ratio is larger than the threshold \(\gamma\) this decides \(\mathcal{H}_1, \hat{\Theta}_1\) is the maximum likelihood estimate (MLE) of the parameters when hypothesis \(\mathcal{H}_1\) is true. In general the GLRT test for the following two hypotheses:

\[
\mathcal{H}_0 : A\Theta = b, \quad \sigma_w > 0 \\
\mathcal{H}_1 : A\Theta \neq b, \quad \sigma_w > 0
\]  

(27)

The linear model is given by \(\mathbf{x} = \mathbf{H}\Theta + \mathbf{w}\). The problem given by equation 25 can be transformed to this problem if we choose \(\Theta = [A], \mathbf{A} = [1]\) and \(b = [0]\). The matrix \(\mathbf{H}\) is chosen to be a column of 1 corresponding to the window size \(N\).

\[
T(x) = (N-1) \frac{(\mathbf{A}\hat{\Theta}_1 - b)^\top (\mathbf{A}\hat{\Theta}_1 - b)}{\mathbf{w}^\top \mathbf{w}}
\]

(28)

Inserting the values of \(\mathbf{A}, \mathbf{b}, \mathbf{H}\) and \(\hat{\Theta}_1\) gives

\[
T(x) = (N - 1) \frac{\hat{x}^2}{\sigma_w^2}
\]

(29)

Where \(\hat{x}\) is the sample mean and the MLE of the variance is given by:

\[
\sigma_w^2 = \frac{1}{N} \mathbf{x}^\top \left( \mathbf{I} - \mathbf{H} \left( \mathbf{H}^\top \mathbf{H} \right)^{-1} \mathbf{H}^\top \right) \mathbf{x}
\]

(30)

which just ends up with being the variance of the test data. The final test statistic is therefore

\[
T(x) = (N - 1) \frac{\hat{x}^2}{\text{var}(x)} > \gamma'
\]

(31)

A more elaborate derivation of the detector can be found in Kay (1998). The probability of false alarms \(P_{FA}\) and the probability of detection \(P_D\) is given by:

\[
P_{FA} = Q_{F_{\gamma'}}(\gamma')
\]

(32)

and

\[
P_D = Q_{F_{\gamma'}}(\gamma') ; \quad \gamma = \frac{N A^2}{\sigma_w^2}
\]

(33)

where \(Q\) denotes the right tail probability.

5.2 Pre-whitened residual detector

The second detector works on the whitened residuals. Since these are distributed according to a Cauchy distribution (as seen from figure 4) the following probability distribution function must be used

\[
p(x; x_o, \beta) = \frac{\beta}{\pi (x - x_o)^2 + \beta^2}
\]

(34)

The two parameters is the half-width half-maximum scaling, \(\beta\), and the offset \(x_o\). The hypotheses are the same as in the nonwhite case, is a DC-level present in the noise or not. The GLRT test statistic becomes

\[
L_G(x) = \frac{\prod_{i=1}^N p((x_i; x_o, \beta))}{\gamma} > \gamma
\]

(35)

The MLEs of \(\beta\) and \(x_o\) is found by fitting the data to equation 34.

The performance can be calculated using these formulas:

\[
P_{FA} = Q_{\chi^2(\lambda)}(\gamma')
\]

(36)

and

\[
P_D = Q_{\chi^2(\lambda)}(\gamma') ; \quad \lambda = \frac{N A^2}{\sigma_w^2}
\]

(37)

Fig. 6. Performance of the two detectors plotted as the difference between probability of detections and false alarms. Note that the value of \(\gamma\) and \(\gamma'\) can not be compared numerically.

Using the performance equations the compare value \(\gamma\) can be found. A plot of the detection performance minus the probability of false alarms for both the detector using the raw residual values and the one working on the whitened residuals are shown. For both performance curves a DC-level of 5 m/s has been chosen as the value wanted to be detected. Difference below this can be down to inaccuracies in the estimation of parameters in the velocity calculations.

6. RESULTS

To test the two detectors, they were exposed to a data record containing a pitot tube defect. The results are shown together with the residuals in Figures 7 and 8.
was designed to pre-whiten the residuals before applying a change detector. A dedicated detector was derived as a generalised likelihood ratio test for the log-likelihood ratio of probability density functions with and without faults. Change detection was obtained with thresholds calculated for the Cauchy distribution that pre-whitened residuals were found to follow. Detection thresholds were determined to find an optimal balance between probabilities for correct detection and false alarms. Two detectors were designed and tested against data records from an UAV with a pitot tube defect that developed while the UAV was in the air. Both detectors were able to detect the fault within an acceptable time frame and demonstrated that a hard landing could have been avoided with the proposed diagnostic system.

7. CONCLUSION

The UAV’s airspeed sensor faults were detected in this paper using a setup where inherent redundancies were used to generate residuals. Information from an onboard GPS receiver and the air speed estimated from engine thrust were used as sources of redundant information. Statistical properties of residuals were investigated and a whitening filter was designed to pre-whiten the residuals before applying a change detector. The UAV’s airspeed sensor faults were detected in this paper using a setup where inherent redundancies were used to generate residuals. Information from an onboard GPS receiver and the air speed estimated from engine thrust were used as sources of redundant information. Statistical properties of residuals were investigated and a whitening filter was designed to pre-whiten the residuals before applying a change detector.