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Infrastructure and Spatial Effects on the Frequency of Cyclist-Motorist Collisions in the Copenhagen Region

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ABSTRACT
Promoting cycling aims at reducing congestion and pollution and encouraging healthy and sustainable lifestyles, but generally clashes with the crash risk perception while riding a bicycle that is still the most significant disincentive to cycling. The current study analyzed the factors contributing to increase crash risk while riding a bicycle by focusing on the variation of 5349 cyclist-motorist collisions within 269 traffic zones in the Copenhagen Region. The model controlled for traffic exposure for both bicycles and motorized transport modes, evaluated the effects of infrastructure and socio-economic characteristics of the zones, and accounted for heterogeneity and spatial correlation across the zones. A Poisson-lognormal model with second-order CAR priors confirmed the existence of the safety in numbers phenomenon, contradicted previous literature about bicycle facilities not being helpful in reducing crash risk, highlighted the need for Copenhagen-style bicycle paths especially in suburban areas, and emphasized how heterogeneity and spatial correlation play a significant role in explaining the probability of cyclist-motorist crash occurrence.
INTRODUCTION

In a world where natural resources are rapidly depleting and health concerns are constantly growing, promoting bicycle use has been embraced by urban and transport planners as a way to encourage a healthy and environmentally sustainable lifestyle (e.g., 1, 2, 3, 4, 5, 6). This lifestyle offers commuters and leisure travelers the possibility of a healthier and more affordable door-to-door transport solution and bears the potential of reducing congestion and pollution that deplete natural resources worldwide (e.g., 7, 8).

Currently, the debate about whether cycling actually provides more benefits or disbenefits is vivid. While some evidence points in the direction of cycling offering advantages because of the physical activity that reduces the health risks inherent to a sedentary life (e.g., 9, 10, 11), other evidence points in the direction of cycling bringing along negative effects because of the exposure to traffic pollution and crash risk (e.g., 12, 13, 14). When considering that the crash risk on a bicycle is definitely higher than the one on a motorized vehicle (e.g., 15, 16, 17) and factoring that it is the most significant disincentive to cycling (e.g., 18, 19, 20), investigating the factors contributing to the risk of being involved in a crash while riding a bicycle assumes a fundamental role for urban and transport planners aiming to increase the appeal of their cities and regions to potential cyclists.

The literature on cycling crashes presents several attempts of understanding the factors contributing to crash occurrence. In particular, a great deal of attention has been posed to infrastructure characteristics and specific conditions leading to a high number of conflicts between cyclists and other road users. Intersections and roundabouts have been often identified as blackspots for cyclists, with particular problems in signalized intersections or roundabouts replacing other types of intersections (e.g., 16, 21, 22, 23). Interruptions to the traffic flow such as the presence of public transport stops or the adoption of traffic calming measures have been generally identified as problematic for cyclists, with specific hindrances linked to the interaction with other road users (e.g., 22, 24, 25). Cycling facilities have been related to higher cycling crash occurrence and these findings have been interpreted as on-road cycling being safer, with specific problems related to bidirectional facilities and discontinuous ones (e.g., 26, 27, 28). Traffic congestion has been analyzed with respect to cycling crash occurrence via surrogate measures such as measuring rates in peak or off-peak hours, considering speed limits, and observing traffic composition, with identifiable problems related both to higher traffic volumes in the peak hours and higher speeds in the off-peak hours (e.g., 29, 30, 31, 32).

Limitations are identifiable in the existing literature on cycling crashes. Firstly, existing research efforts generally examined the correlation of crash occurrence with one specific factor and hence missed on the multifaceted and interrelated nature of the contribution of different factors to the number of cycling crashes. Secondly, existing research endeavors largely overlooked the heterogeneity of the zone or road where the crashes happened as well as the spatial correlation effects that evidently exist as crashes are more frequent in certain locations. Thirdly, existing research studies ignored the exposure in terms of traffic and hence missed on correcting for the amount of cyclists in the zone or road where the crashes took place while likely introducing bias in the estimates of the effects related to crash occurrence. The current study overcomes these limitations in the existing literature by proposing a comprehensive analysis of the frequency of cyclist-motorist crashes that considers heterogeneity and spatial correlation across zones.

The current study estimated a crash frequency model of 5349 cyclist-motorist collisions occurred in the Copenhagen Region between 2009 and 2013. Crash data were
available from the National Crash Database maintained by the Danish Road Directorate and compiled from police records, and their geographical coordinates were matched to the road network of the study region. The model formulation considered traffic zones within the region as the level of aggregation of the crash counts and accounted for both heterogeneity and spatial correlation across the zones. The model specification focused on the infrastructure and the characteristics of the traffic zones in the region, but most relevantly corrected for the traffic exposure of bicycles, cars, vans and heavy vehicles. In fact, the availability of a large dataset of actual travel behavior from the on-going national travel survey and the accessibility to traffic estimates from regional and national transport models allowed overcoming the main limitation of crash occurrence studies not correcting for traffic exposure. The estimation of Poisson-gamma and Poisson-lognormal models with conditional autoregressive (CAR) priors within a full hierarchical Bayesian framework allowed answering the requirements for the formulation and the specification of the model.

Analyzing cyclist-motorist crashes in the Copenhagen Region has several advantages. Firstly, it allows providing an outer marker for emphasizing safety issues and suggesting traffic and policy measures that are relevant for the realization of cycling safety in Denmark as well as in other countries (see, e.g., 33). Established cycling cultures could find interesting looking at the Danish experience and evaluating similarities and differences while estimating a similar model. Emerging cycling cultures could find inspiration in the Danish example and reflect upon policies bearing the potential to reduce the probability of cyclist-motorist crashes. Secondly, it allows suggesting the type of information that is necessary for analyzing cyclist-motorist collisions and properly designing policies with the aim of promoting bicycle use, while providing the availability of detailed information about exposure and hence avoiding to obtain biased estimates that might lead to counterintuitive results. Thirdly, it answers a specific need for a Danish society where the safety of cyclists is at the top of the priority list when considering cycling strategies being designed and implemented at the municipal, regional and national level.

The remainder of this paper is structured as follows. The next section describes the data collection effort and introduces the model formulation and specification. Then, model estimates are presented and discussed with respect to findings from the existing literature. Last, conclusions are drawn and further research directions are proposed.

METHODS

Data

The current study focused on 5349 cyclist-motorist collisions that occurred in the Copenhagen Region from 2009 to 2013. Crashes were extracted from the National Crash Database that the Danish Road Directorate compiles from police reports. These reports enclose information about the characteristics of the crash, the injured persons and the involved vehicles (for details, see, e.g., 34, 35). Most importantly, these reports contain a georeferenced location of the crashes that enables matching them with the road network.

The selection of collisions between cyclists and motorists was motivated by the possible bias that would have been introduced by considering also single-bicycle crashes or collisions between cyclists and vulnerable road users that are subject to a high degree of underreporting (see, e.g., 12, 36, 37). Collisions between cyclists and motorists are far more reported also because insurance claims require the police to be called. The choice of a five-year period was induced by the need for having a long enough period to assemble a sample of adequate size and a short enough period to limit changes in road and traffic conditions. The
focus of the study on the Copenhagen Region was prompted by the need for a very detailed road network and a calibrated transport model for planning purposes in order to be able to consider infrastructure characteristics and traffic volumes within the model specification.

The geocoded crashes were matched to the bicycle network of the Copenhagen Region. This is a high-resolution network of 110,893 nodes and 272,586 links that was constructed from a variety of sources for planning purposes. The network contains paths only accessible by bicycles and pedestrians, bicycle lanes and paths alongside roads for motorized traffic, bicycle paths along motorways and expressways, and all the road network available to motorists in the region. The network covers an area of about 3000 sq.km. and a population of about 2 million inhabitants divided in 18 municipalities and 269 traffic zones according to the zone system of the Danish national transport model (LandsTrafikModel, LTM). The information about the traffic on the network was obtained by working on three sources: the national travel survey (TransportvaneUndersøgelse, TU), the regional transport model (ØresundTrafikModel, OTM), and the LTM. The TU is an on-going collection of 1000 travel diaries per month of a representative sample of the Danish population between 10 and 84 years old. The OTM is the transport planning model traditionally used for traffic evaluations in the Copenhagen Region. The LTM is the transport planning model that will be used for traffic and transport policy assessments in Denmark. Both the OTM and the LTM use the TU travel diaries for the construction of the origin-destination matrices and the estimation of the demand model. The current study used the OTM matrices for cars, vans, heavy vehicles and bicycles in order to assign the inter-zonal traffic on the network, and enriched the information with the TU travel diaries in order to assign the bicycle intra-zonal traffic on the network. In fact, considering intra-zonal traffic was very relevant to properly assess the bicycle traffic as traditionally the distances covered by bicycle are lower than the ones traveled by motorized transport modes. The current study then used the traffic zones from the LTM because of the high level of detail about the zones in which the Copenhagen Region is divided.

In light of the described data availability, the variables for the model specification were constructed as follows: (i) crash data were considered as count variables representing the number of crashes in a spatial unit, and this unit corresponded to the LTM traffic zones; (ii) infrastructure data for each LTM traffic zone were obtained by overlaying the network data with the traffic zone boundaries; (iii) traffic exposure was retrieved by taking the average daily traffic (vehicle \( \times \) km for cars, vans, heavy vehicles and bicycles) at the link level and then aggregating the values at the LTM traffic zone level. While a limitation of the current study might consist in considering the average daily traffic over the five-year period, it is offset by the benefit of considering traffic exposure within crash frequency models. Notably, the traffic exposure was validated against a large database of traffic counts and the variations in the counts over the period were not conspicuous, especially for the cycling traffic that has not shown significant alterations in the period between 2000 and 2011 (38).

Table 1 presents the most significant variables in the dataset for model estimation. This study focused on the Copenhagen Region, where the LTM traffic zones are quite heterogeneous in terms of area and population, with smaller zones in the Copenhagen city center and larger zones in the outskirts of the region. The distribution of the average income does not reflect completely one of the jobs, as higher income is found in the north of the metropolitan area while most jobs are found in the Copenhagen city center (see, e.g., 35). Notably, the development of job opportunities and the expansion of the metropolitan area has followed the transit-oriented fingerplan and accordingly has been around the five train lines from the center of the city (see, e.g., 35). The Copenhagen Region has an extremely
developed bicycle infrastructure, especially in the Copenhagen city center and the residential zones that have adopted the typical Copenhagen-style design for bicycle paths (i.e., road-curb-bicycle path-curb-sidewalk). This design has been found to increase the perception of safety with respect to alternative mixed designs and hence to be at the root of the high bicycle market share \((39)\). Cyclists share the road though with motorists, in particular in the outskirts of the metropolitan area and the suburban areas of the region.

### TABLE 1 Sample Characteristics

<table>
<thead>
<tr>
<th>Variable</th>
<th>mean</th>
<th>st.dev.</th>
<th>min</th>
<th>25% pct.</th>
<th>median</th>
<th>75% pct.</th>
<th>max</th>
</tr>
</thead>
<tbody>
<tr>
<td>Number of crashes</td>
<td>19.9</td>
<td>22.7</td>
<td>0</td>
<td>6</td>
<td>13</td>
<td>26</td>
<td>194</td>
</tr>
<tr>
<td>Population (unit)</td>
<td>6867.7</td>
<td>4624.1</td>
<td>0</td>
<td>3590</td>
<td>6059</td>
<td>9472</td>
<td>30190</td>
</tr>
<tr>
<td>Area (sq.km.)</td>
<td>11.5</td>
<td>18.0</td>
<td>0.161</td>
<td>1.391</td>
<td>4.25</td>
<td>13.982</td>
<td>123.579</td>
</tr>
<tr>
<td>Average income (1000 DKK)</td>
<td>228.8</td>
<td>50.5</td>
<td>85.5</td>
<td>199.3</td>
<td>220.2</td>
<td>249</td>
<td>501.2</td>
</tr>
<tr>
<td>Full-time employed (unit)</td>
<td>3931</td>
<td>2784</td>
<td>0</td>
<td>2025</td>
<td>3468</td>
<td>5416</td>
<td>18342</td>
</tr>
<tr>
<td>Part-time employed (unit)</td>
<td>163</td>
<td>125</td>
<td>0</td>
<td>87</td>
<td>138</td>
<td>194</td>
<td>822</td>
</tr>
<tr>
<td>Students or pupils (unit)</td>
<td>1263</td>
<td>837</td>
<td>0</td>
<td>688</td>
<td>1128</td>
<td>1769</td>
<td>5435</td>
</tr>
<tr>
<td>Retired (unit)</td>
<td>1139</td>
<td>737</td>
<td>0</td>
<td>535</td>
<td>1066</td>
<td>1622</td>
<td>3104</td>
</tr>
<tr>
<td>Unemployed seeking job (unit)</td>
<td>117</td>
<td>160</td>
<td>0</td>
<td>25</td>
<td>59</td>
<td>144</td>
<td>1092</td>
</tr>
<tr>
<td>Unemployed not seeking job (unit)</td>
<td>255</td>
<td>227</td>
<td>0</td>
<td>96</td>
<td>200</td>
<td>348</td>
<td>1577</td>
</tr>
<tr>
<td>Road without bicycle lane (km)</td>
<td>58.760</td>
<td>58.378</td>
<td>1.025</td>
<td>16.996</td>
<td>44.567</td>
<td>79.875</td>
<td>403.113</td>
</tr>
<tr>
<td>Road with bicycle lane (km)</td>
<td>6.496</td>
<td>5.103</td>
<td>0.798</td>
<td>3.120</td>
<td>5.894</td>
<td>8.860</td>
<td>31.600</td>
</tr>
<tr>
<td>Road with bicycle path (km)</td>
<td>22.511</td>
<td>21.570</td>
<td>1.241</td>
<td>4.905</td>
<td>18.043</td>
<td>32.518</td>
<td>135.831</td>
</tr>
<tr>
<td>Footpath segregated (km)</td>
<td>2.217</td>
<td>2.574</td>
<td>0.162</td>
<td>0.453</td>
<td>1.427</td>
<td>2.828</td>
<td>14.523</td>
</tr>
<tr>
<td>Cycling traffic (bicycle × km)</td>
<td>8.633</td>
<td>6.656</td>
<td>137</td>
<td>4.349</td>
<td>7.117</td>
<td>11.281</td>
<td>62.431</td>
</tr>
<tr>
<td>Car traffic (vehicle × km)</td>
<td>101,489</td>
<td>101,750</td>
<td>4.438</td>
<td>34,820</td>
<td>62,181</td>
<td>135,637</td>
<td>523,433</td>
</tr>
<tr>
<td>Van traffic (vehicle × km)</td>
<td>10,316</td>
<td>11,010</td>
<td>477</td>
<td>3,217</td>
<td>6,113</td>
<td>14,092</td>
<td>62,341</td>
</tr>
<tr>
<td>Heavy vehicle traffic (vehicle × km)</td>
<td>5,049</td>
<td>5,891</td>
<td>54</td>
<td>1,355</td>
<td>2,909</td>
<td>6,587</td>
<td>32,197</td>
</tr>
</tbody>
</table>

### Model

The current study intended to analyze the frequency of collisions between cyclists and motorists by estimating a count data model at the zone level while considering heterogeneity and spatial correlation across the zones. Accordingly, Poisson-based models with heterogeneity effects and conditional autoregressive priors were estimated as recommended in the literature (for a review, see 40, 41).

Considering the crash counts within the traffic zones in the Copenhagen Region, the base form of the Poisson-based model was expressed as:

\[
Y_i \sim \text{Poisson}(\mu_i)
\]

\[
\log(\mu_i) = \alpha + \beta X_i + v_i + u_i
\]

where \(Y_i\) is the observed number of crashes in traffic zone \(i\), \(\mu_i\) is the expected Poisson crash rate in traffic zone \(i\), \(X_i\) is a vector of explanatory variables, \(\alpha\) is the intercept to be estimated,
\( \beta \) is a vector of parameters to be estimated, \( v_i \) is a random term that captures the heterogeneity across traffic zones, and \( u_i \) is a random term that captures the spatial correlation across traffic zones. The vector \( X_i \) of explanatory variables includes the traffic exposure (i.e., cars, vans, trucks, bicycles), the infrastructure characteristics and the zone characteristics. Different model formulations were considered according to different specifications of the error terms \( v_i \) and \( u_i \): Poisson-lognormal and Poisson-gamma models were tested to evaluate which distribution of the random term \( v_i \) was the most suitable to account for heterogeneity across zones, while two neighboring structures were tested to assess which structure expressed in the random term \( u_i \) was the most appropriate to account for spatial correlation across zones.

Both the Poisson-lognormal and the Poisson-gamma formulations assigned a uniform prior distribution to the intercept \( \alpha \) and a highly non-informative normal prior to all \( \beta \)'s with zero mean and 100,000 variance (see, e.g., 42). In the Poisson-lognormal formulation, the prior distribution for the heterogeneity random term \( v_i \) was a normal prior with distribution \( N(0, \tau_v^{-1}) \), where \( \tau_v \) is the precision (i.e., the inverse of the variance) with a vague gamma prior Gamma(0.5, 0.001). It should be noted that the parameterization of the Gamma distribution Gamma(\( a, b \)) has mean \( ab \) and variance \( ab^2 \). In the Poisson-gamma formulation, the prior distribution for the exponential of the heterogeneity random term \( \exp(v_i) \) was a gamma prior with distribution Gamma(\( \phi, \phi \)) where \( \phi \) was assigned to a non-vague hyper prior with Gamma(0.1, 1.0).

Both formulations for the correlation structure across zones considered the same priors for the intercept \( \alpha \) and the parameters \( \beta \)'s, and both considered the two formulations of the heterogeneity random term \( v_i \). The spatial correlation term \( u_i \) was represented with a conditional autoregressive model (43):

\[
  u_i | u_j, j \neq i \sim N \left( \sum_j u_j w_{ij}, \frac{\tau_{ij}^2}{w_{ij}} \right)
\]

(3)

where \( w_{ij} \) is the weight between zone \( i \) and zone \( j \), \( w_{ij} \) is the sum over \( j \) of the weights \( w_{ij} \), and \( \tau_{ij}^2 \) is a scale parameter with a prior with distribution Gamma(0.5, 0.001). The two formulations of the correlation structure were differentiated according to the definition of the weights \( w_{ij} \). In the first formulation, first-order neighbours were defined for traffic zone \( j \) sharing a border with traffic zone \( i \) with \( w_{ij} = 1 \), and for traffic zones not sharing any border with \( w_{ij} = 0 \). In the second formulation, second-order neighbours were also defined for traffic zone \( j \) being connected to first-order neighbors of zone \( i \) with \( w_{ij} = 0.5 \), and with \( w_{ij} = 0 \) otherwise.

Models were estimated with the software package Openbugs (44) by using the Markov Chain Monte Carlo (MCMC) method under the full hierarchical Bayesian framework. As several models were estimated (i.e., the two heterogeneity formulations without spatial correlation, and the four combinations of the formulations for the error terms \( v_i \) and \( u_i \)), the deviance information criterion (DIC) was used to compare the goodness-of-fit and select the best model (42).

RESULTS

The specification of the model followed a traditional iterative process looking for the best expression of variables and their transformation in order to explain the variation of the number of collisions between cyclists and motorists in the Copenhagen Region. All the six
model formulations were then estimated with the same specification of variables and parameters throughout the iterative process.

The average daily traffic values for bicycles, cars, vans, and heavy vehicles, were transformed in logarithmic scales in order to reduce the large variation of these explanatory variables across zones. The same applied to the number of kilometers of road with and without cycling facilities, whose effect was investigated also relatively to their location in either urban or suburban areas of the Copenhagen Region. The socio-economic characteristics of the zones were also considered in the model specification in order to include proxies relative to the availability of resources for infrastructure development and maintenance, as the income in a zone reflects the taxation resources of the local authorities and the typology of demand given that the profile of the population in a zone reflects the potential heterogeneity in the demand composition. The various specifications were tested with all six model formulations described in the methodological section, and the comparison of the DIC revealed that Poisson-lognormal models performed better than Poisson-gamma models, and that second-order spatial correlation effects led to better model fit than first-order effects only. Table 2 presents the comparison of the DIC for the six model formulations with the best model specification, with differences that are significant according to the indications by Spiegelhalter et al. (42).

**TABLE 2 Comparison of Model Formulations**

<table>
<thead>
<tr>
<th>Model</th>
<th>Heterogeneity</th>
<th>Spatial correlation</th>
<th>DIC</th>
</tr>
</thead>
<tbody>
<tr>
<td>Aspatial Poisson-lognormal</td>
<td>Lognormal</td>
<td>-</td>
<td>1345.7</td>
</tr>
<tr>
<td>Aspatial Poisson-gamma</td>
<td>Gamma</td>
<td>-</td>
<td>1361.8</td>
</tr>
<tr>
<td>Spatial Poisson-lognormal</td>
<td>Lognormal</td>
<td>1st order CAR</td>
<td>1325.5</td>
</tr>
<tr>
<td>Spatial Poisson-gamma</td>
<td>Gamma</td>
<td>1st order CAR</td>
<td>1337.1</td>
</tr>
<tr>
<td>Spatial Poisson-lognormal</td>
<td>Lognormal</td>
<td>2nd order CAR</td>
<td>1305.9</td>
</tr>
<tr>
<td>Spatial Poisson-gamma</td>
<td>Gamma</td>
<td>2nd order CAR</td>
<td>1315.8</td>
</tr>
</tbody>
</table>

The posterior means and standard deviations of the intercept $\alpha$, the $\beta$’s for the explanatory variables, the standard deviation of the heterogeneity error term $\nu_i$ and the standard deviation of the spatial correlation error term $u_i$ were estimated using the MCMC method. For all the model formulations and specifications, two chains were simulated with different initial values and the initial 25,000 iterations were discarded as burn-ins to reach the convergence of the two chains. Then, both chains were simulated for other 75,000 iterations with the aim of calculating the posterior means and standard deviations of the estimated parameters for the best model, namely the Poisson-lognormal with second-order CAR priors, that are presented in table 3.

The average daily traffic is related to the number of collisions between cyclists and motorists in each zone. Interestingly, the estimated parameter shows a non-linear relationship between the number of crashes and the bicycle × km with a statistically significant difference from 1, as credible sets do not contain 1 neither at the 95% nor at the 90% confidence level. This is an important result because it shows that, ceteris paribus, the crash rates fall as cycling exposure increases, most likely because higher numbers of cyclists increase awareness in drivers and hence reduce risk. Also, the estimated parameters show a non-linear relationship between the number of crashes and the vehicle × km for all the types of motorized traffic.
considered, namely cars, vans and heavy vehicles. This is also a relevant result because it suggests that the crash rates decrease with increasing volumes of traffic, or in other words with increasing congestion.

### TABLE 3 Poisson-lognormal CAR model estimates

<table>
<thead>
<tr>
<th>Variable</th>
<th>Mean</th>
<th>st.dev.</th>
<th>MC err.</th>
<th>sig.</th>
</tr>
</thead>
<tbody>
<tr>
<td>log (bicycle × km)</td>
<td>0.6690</td>
<td>0.0361</td>
<td>0.0021</td>
<td>**</td>
</tr>
<tr>
<td>log (car × km)</td>
<td>0.3490</td>
<td>0.0877</td>
<td>0.0052</td>
<td>**</td>
</tr>
<tr>
<td>log (van × km)</td>
<td>0.3150</td>
<td>0.1097</td>
<td>0.0065</td>
<td>**</td>
</tr>
<tr>
<td>log (heavy vehicle × km)</td>
<td>0.1746</td>
<td>0.0812</td>
<td>0.0048</td>
<td>**</td>
</tr>
<tr>
<td>Copenhagen/Frederiksberg</td>
<td>-0.0492</td>
<td>0.0447</td>
<td>0.0019</td>
<td></td>
</tr>
<tr>
<td>Suburban areas</td>
<td>0.6141</td>
<td>0.2039</td>
<td>0.0093</td>
<td></td>
</tr>
<tr>
<td>log (km road without bicycle lane)</td>
<td>0.2755</td>
<td>0.0540</td>
<td>0.0019</td>
<td>**</td>
</tr>
<tr>
<td>log (km road with bicycle lane)</td>
<td>0.0565</td>
<td>0.0247</td>
<td>0.0002</td>
<td>**</td>
</tr>
<tr>
<td>log (km road with bicycle path)</td>
<td>-0.2928</td>
<td>0.0642</td>
<td>0.0016</td>
<td>**</td>
</tr>
<tr>
<td>log (km road without bicycle lane) in suburban areas</td>
<td>0.0412</td>
<td>0.0260</td>
<td>0.0003</td>
<td></td>
</tr>
<tr>
<td>log (km road with bicycle lane) in suburban areas</td>
<td>-0.1230</td>
<td>0.1309</td>
<td>0.0059</td>
<td>**</td>
</tr>
<tr>
<td>log (km road with bicycle path) in suburban areas</td>
<td>0.1555</td>
<td>0.0644</td>
<td>0.0006</td>
<td></td>
</tr>
<tr>
<td>number of intersections</td>
<td>0.0414</td>
<td>0.0188</td>
<td>0.0003</td>
<td>**</td>
</tr>
<tr>
<td>number of intersections in suburban areas</td>
<td>-0.0246</td>
<td>0.0116</td>
<td>0.0002</td>
<td>**</td>
</tr>
<tr>
<td>log (income)</td>
<td>-0.1499</td>
<td>0.0989</td>
<td>0.0059</td>
<td>*</td>
</tr>
<tr>
<td>full-time employed</td>
<td>1.1385</td>
<td>0.3903</td>
<td>0.0231</td>
<td>**</td>
</tr>
<tr>
<td>part-time employed</td>
<td>0.0549</td>
<td>0.0306</td>
<td>0.0018</td>
<td>*</td>
</tr>
<tr>
<td>students or pupils</td>
<td>-0.4595</td>
<td>0.0902</td>
<td>0.0053</td>
<td>**</td>
</tr>
<tr>
<td>Retired</td>
<td>0.1788</td>
<td>0.0604</td>
<td>0.0035</td>
<td>**</td>
</tr>
<tr>
<td>population on welfare</td>
<td>0.8984</td>
<td>0.3201</td>
<td>0.0186</td>
<td>**</td>
</tr>
<tr>
<td>Constant</td>
<td>-2.5614</td>
<td>1.1020</td>
<td>0.0657</td>
<td>**</td>
</tr>
<tr>
<td>st.dev. (v)</td>
<td>0.3160</td>
<td>0.0554</td>
<td>0.0018</td>
<td>**</td>
</tr>
<tr>
<td>st.dev. (u)</td>
<td>0.2622</td>
<td>0.0596</td>
<td>0.0020</td>
<td>**</td>
</tr>
</tbody>
</table>

* statistically significant difference from zero (90% credible set shows the same sign)
** statistically significant difference from zero (95% credible set shows the same sign)

mean: mean of the distribution of the estimated parameter; st.dev.: standard deviation of the distribution of the estimated parameter; MC err: Monte Carlo error

After controlling for the exposure, the location of the zone is significantly associated with the number of crashes, as more crashes are related to suburban areas. Most relevantly, a higher extension of bicycle facilities is significantly related to a decrease in the number of crashes, in particular when longer bicycle paths that are segregated from the vehicle traffic are constructed. The same effect is not observed for bicycle lanes alongside the roads, although
not as remarkably as for the absence of bicycle facilities, thus confirming that the Copenhagen-style bicycle path design is the safest solution to reduce the risk while cycling. Notably, when considering the interaction with the location in urban or suburban areas, it emerges that bicycle paths are less effective in suburban areas. When looking at the socio-economic characteristics of each zone, there is a negative relationship between the number of crashes and the average income, suggesting that a lower number of crashes occurs in zones that are in average richer and hence have higher taxation intake. The number of crashes is also associated with a higher number of potential commuters, identifiable with the number of full-time and part-time workers, a lower number of students, and a higher number of people being maintained by the welfare system.

When looking at the heterogeneity and spatial correlation, the posterior mean of the standard deviation of the error term accounting for heterogeneity is statistically significant and suggests that heterogeneity effects are relevant in these crash data. Analogously, the posterior mean of the standard deviation of the error term representing the spatial correlation is statistically significant and indicates that spatial correlation effects play a significant role in these crash data. It should be noted that the best model specification has spatial effects concerning not only the obvious correlation with zones that are direct neighbors, but also the more loose correlation with zones sharing a common neighbor. The ratio between the st.dev. (\(\mu\)) of the error term representing the spatial correlation and the sum of the standard deviations st.dev. (\(\mu\)) and st.dev. (\(\nu\)) of both error terms allows assessing the posterior proportion of model error explained by spatial random effects. The value of this ratio indicates that the spatial correlation accounts for 54.7% of the stochastic variation in the model and hence assumes a slightly more prominent role with respect to the heterogeneity.

**DISCUSSION AND CONCLUSIONS**

The current study analyzed the frequency of cyclist-motorist collisions in the Copenhagen Region by estimating zone-based Poisson-based models accounting for heterogeneity and spatial correlation across zones. In particular, model estimates illustrate the importance of infrastructure effects and the relevance of spatial correlation to explain the variation in the number of crashes within the region studied. The best model for these crash data is a Poisson-lognormal with second-order CAR priors that account for correlation not only between neighboring zones, but also between zones sharing a neighbor.

An extremely relevant finding is the non-linear relationship between crashes and average bicycle daily traffic in the zone. As the traffic is expressed as bicycle \(\times\) km and the parameter is significantly lower than 1, this non-linearity confirms the safety in numbers hypothesis that the crash rates diminish when the bicycle traffic increases (e.g., 1, 3). Having a higher number of cyclists can be motivated from both the cyclists’ perspective, namely they feel safer in an environment designed for them that they share with a lot of fellow cyclists, and the drivers’ perspective, namely they reach higher awareness because they see cyclists everywhere. Non-linearity is observed also with the vehicle traffic, and again as the traffic is expressed as vehicle \(\times\) km and the parameter is significantly lower than 1, this non-linearity suggests that the crash rates diminish when vehicular traffic augments. This finding is not surprising, especially when considering that congestion has been generally related to lower crash rates because of the lower speeds and the consequent higher margins of error in recognizing a potential conflict (e.g., 16, 18, 23).

Another extremely important finding is the positive correlation between the presence of bicycle infrastructure and the reduction in the number of crashes. Findings from the current
study are in disagreement with previous literature showing that cycling in mixed traffic is safer than cycling on bicycle infrastructure (e.g., 26, 27, 28). These findings are however in line with recent studies pointing out that bicycle infrastructure not only increases safety (35, 45), but also increases the perception of safety that inherently encourages more people to cycle (39). It should be noted that the estimated model controls for traffic exposure and for a very detailed representation of the extensive infrastructure in one of the cycling capital regions of the world, while previous studies from the 90’s were from regions with very limited infrastructure and very little bicycle traffic that was not controlled for. Accordingly, non-significant relationships between crash rates and bicycle infrastructure could have resulted from lack of observations rather than actual absence of correlation. It should also be noted that the findings of the current study extend the discussion by underlying the importance of the infrastructure in suburban areas where the separation between cyclists and motorists is even more beneficial because of the risk inherent in the higher speeds of the vehicles (e.g., 31, 32). Findings about intersections being problematic are also in general in line with previous research (16, 21, 22), and also in this case they extend the current knowledge by emphasizing in particular that the effect is less marked in suburban areas than in urban ones.

The relevance of heterogeneity and spatial correlation is evident from the model estimates, as is the relative importance that shows how the variation of the error term is slightly towards the spatial effects. This confirms the importance of considering spatial correlation within frequency models (see, e.g., 40, 41) and looking for specific solutions in specific locations where more crashes occur in order to achieve a substantial reduction in crash occurrence. Findings from this study underline how promoting bicycle use can create a virtuous circle in which the more people bike and the safer are the roads for cyclists, how the Copenhagen-style bicycle path design is the most effective for achieving a decrease in the number of crashes and hence an increase in encouraging cycling, and how intersection solutions are necessary to reduce the number of conflicts. Future research should investigate whether the possibility of a model estimated at the link level is feasible, and whether a joint model of frequency and severity could help further understanding the determinants of cyclist-motorist collisions.

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REFERENCES


